

Stationary Processes

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Overview

1. Basic Properties and Linear Processes
2. Introduction to ARMA Process
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AutoCovariance(ACVF) and AutoCorrelation(ACF)

$$\text{ACVF: } \gamma(h) = \text{Cov}(X_{t+h}, X_t)$$

$$\text{ACF: } \rho(h) = \frac{\gamma(h)}{\gamma(0)}$$

$$h = 0, \pm 1, \pm 2, \dots$$

ACVF and ACF provide a useful measure of the degree of dependence among the values of time series at different times.

Example

Problem: Suppose X_t is a stationary Gaussian* time series and we observed X_n . Find a function of X_n that gives us the best predictor of X_{n+h} .

Solution: The conditional distribution of X_{n+h} given that $X_n = x_n$ is $N(\mu + \rho(h)(x_n - \mu), \sigma^2(1 - \rho(h)^2))$, where μ and σ^2 are mean and variance of X_t .

$$X_{n+h} = m(X_n) = \mu + \rho(h)(X_n - \mu)$$

The corresponding mean square error is

$$E(X_{n+h} - m(X_n))^2 = \sigma^2(1 - \rho(h)^2)$$

* X_t is a **Gaussian time series** if all of its joint distributions are multivariate normal, i.e., if for any collection of integers i_1, \dots, i_n the random vector $(X_{i_1}, \dots, X_{i_n})'$ has a multivariate normal distribution

Example(cont')

Remark 1: Prediction of X_{n+h} in terms of X_n is more accurate as $|\rho(h)|$ becomes closer to 1, and as $\rho \rightarrow \pm 1$ the best predictor approaches $\mu \pm (X_n - \mu)$ and mean square error approaches 0.

Remark 2: The best linear predictor depends only on the mean and ACf of series X_t .

Properties of ACVF

- $\gamma(0) \geq 0$
- $|\gamma(h)| \leq \gamma(0)$ for all h
- $\gamma(h) = \gamma(-h)$ for all h

Properties of ACVF(cont')

Definition: A real-valued function k defined on the integers is **nonnegative definite** if

$$\sum_{i,j=1}^n a_i k(i-j) a_j \geq 0$$

for all positive integers n and vectors $\mathbf{a} = (a_1, \dots, a_n)'$ with real-valued components a_i .

Theorem: A real-valued function defined on the integers is the autocovariance function of a stationary time series if and only if it is even and nonnegative definite.

Remark: An autocorrelation function $\rho(\cdot)$ has all properties of an autocovariance function and satisfies the additional condition $\rho(0) = 1$.

Weakly stationary time series

X_t is a **weakly stationary** if

- $\mu_X(t)$ is independent of t
- $\gamma_X(t+h, t)$ is independent of t for each h

Strictly stationary time series (SSTS)

X_t is a **strictly stationary time series** if

$$(X_1, X_2, \dots, X_n)' = (X_{1+h}, X_{2+h}, \dots, X_{n+h})'$$

, for all integers h and $n \geq 1$

A strictly stationary sequence is one in which the joint distribution of these two vectors (and not just the means and covariances) are the same.

Properties of SSTS

- The random variables X_t are identically distributed
- $(X_t, X_{t+h})' = (X_1, X_{1+h})'$ for all integers t and h
- X_t is weakly stationary if $E(X_t^2) < \infty$ for all t
- Weak stationarity does not imply strict stationarity
- An *iid* sequence is strictly stationary

Constructing SSTs

One way to construct a time series X_t that is strictly stationary is to "filter" an *iid* sequence of random variables Z_t and define

$$X_t = g(Z_t, Z_{t-1}, \dots, Z_{t-q})$$

for some real-valued function $g()$. The X_t is **q-dependent**, i.e. that X_s and X_t are independent whenever $|t - s| < q$. (An *iid* sequence is 0-dependent). In the same way a stationary time series is **q-correlated** if $\gamma(h) = 0$ whenever $|h| > q$

The MA(q) Process:

X_t in a **moving-average process of order q** if

$$X_t = Z_t + \theta_1 Z_{t-1} + \dots + \theta_q Z_{t-q},$$

where $Z_t \sim WN(0, \sigma^2)$ and $\theta_1, \dots, \theta_q$ are constants

If X_t is a stationary q-correlated time series with mean 0, then it can be represented as the MA(q) process.

Linear process

The time series X_t in a **linear process** if it has the representation

$$X_t = \sum_{j=-\infty}^{\infty} \psi_j Z_{t-j},$$

for all t , where $Z_t \sim WN(0, \sigma^2)$ and ψ_j is a sequence of constants with $\sum_{j=-\infty}^{\infty} |\psi_j| < \infty$

A linear process is called a **moving average** or **MA(∞)** if $\psi_j = 0$ for all $j < 0$, i.e., if

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j},$$

Introduction to ARMA Processes

The time series X_t in an **ARMA(1,1) process** if it is stationary and satisfies (for every t)

$$X_t - \phi X_{t-1} = Z_t + \theta Z_{t-1},$$

where $Z_t \sim WN(0, \sigma^2)$ and $\phi + \theta \neq 0$

Properties of ARMA

- A stationary solution of the ARMA(1,1) equations exists if and only if $\phi \neq \pm 1$.
- if $|\phi| < 1$ then the unique stationary solution is given by $X_t = Z_t + (\phi + \theta) \sum_{j=1}^{\infty} \phi^{j-1} Z_{t-j}$. In this case we say that X_t is causal or a casual function of Z_t since X_t can be expressed in terms of the current and past values $Z_s, s \leq t$.
- if $|\phi| > 1$ then the unique stationary solution is given by $X_t = -\theta\phi^{-1}Z_t - (\phi + \theta) \sum_{j=1}^{\infty} \phi^{-j-1} Z_{t+j}$. The solution is noncasual, since X_t is then a function of $Z_s, s \geq t$.

Estimation of Mean

$$\bar{X}_n = n^{-1}(X_1 + X_2 + \dots + X_n)$$

If X_t is a stationary time series with mean μ and $\gamma(\cdot)$, then as $n \rightarrow \infty$,

$$\text{Var}(\bar{X}_n) = E(\bar{X}_n - \mu)^2 \rightarrow 0 \text{ if } \mu(n) \rightarrow 0,$$

$$nE(\bar{X}_n - \mu)^2 \rightarrow \sum_{|h|} \mu(h) \text{ if } \sum_{h=-\infty}^{\infty} |\mu(h)| < \infty$$

Estimation of $\mu(\cdot)$ and $\rho(\cdot)$

$$\hat{\gamma}(h) = n^{-1} \sum_{t=1}^{n-|h|} (X_{t+|h|} - \bar{X}_n)(X_t - \bar{X}_n)$$

and

$$\hat{\rho}(h) = \frac{\hat{\gamma}(h)}{\hat{\gamma}(0)}$$

Estimation of $\mu(\cdot)$ and $\rho(\cdot)$ (cont')

The sample ACVF has the desirable property that for each $k \geq 1$ the k -dimensional sample covariance matrix

$$\hat{\Gamma}_k = \begin{bmatrix} \hat{\gamma}(0) & \hat{\gamma}(1) & \dots & \hat{\gamma}(k-1) \\ \hat{\gamma}(1) & \hat{\gamma}(0) & \dots & \hat{\gamma}(k-1) \\ \cdot & \cdot & \dots & \cdot \\ \cdot & \cdot & \dots & \cdot \\ \hat{\gamma}(k-1) & \hat{\gamma}(k-2) & \dots & \hat{\gamma}(0) \end{bmatrix}$$

is nonnegative definite.

Forecasting stationary time series

Predicting the values X_{n+h} , $h > 0$, of a stationary time series with known mean μ and autocovariance γ in terms of values X_n, \dots, X_1 , upto time n .

$$P_n X_{n+h} = a_0 + a_1 X_n + a_n X_1$$

Needed to determine the coefficients a_0, a_1, \dots, a_n , by finding the values that minimizes

$$S(a_0, \dots, a_n) = E(X_{n+h} - a_0 - a_1 X_n - \dots - a_n X_1)^2$$

Forecasting stationary time series (cont')

$$\frac{\partial S(a_0, \dots, a_n)}{\partial a_j} = 0, j = 0, \dots, n$$

Evaluation of derivatives gives the equivalent equations

$$E[X_{n+h} - a_0 - \sum_{i=1}^n a_i X_{n+1-i}] = 0,$$

$$E[(X_{n+h} - a_0 - \sum_{i=1}^n a_i X_{n+1-i})X_{n+1-j}] = 0, j = 1, \dots, n$$

Remark: These 2 last equations determine $P_n X_{n+h}$ uniquely

Forecasting stationary time series(cont')

$$a_0 = \mu \left(1 - \sum_{i=1}^n a_i \right)$$

and

$$\Gamma_n \mathbf{a}_n = \gamma_n(h) \quad (1)$$

where

$$\mathbf{a}_n = (a_1, \dots, a_n)'$$

$$\Gamma_n = [\gamma(i-j)]_{i,j=1}^n$$

$$\text{and } \gamma_n(h) = (\gamma(h), \gamma(h+1), \dots, \gamma(h+n-1))'$$

Forecasting stationary time series (cont')

$$P_n X_{n+h} = \mu + \sum_{i=1}^n a_i (X_{n+1-i} - \mu) \quad (2)$$

The mean square prediction error is

$$E(X_{n+h} - P_n X_{n+h})^2 = \gamma(0) - 2 \sum_{i=1}^n a_i \gamma(h+i-1) + \sum_{i=1}^n \sum_{j=1}^n a_i \gamma(i-j) a_j =$$
$$\gamma(0) - \mathbf{a}'_n \boldsymbol{\gamma}_n(h)$$

Example

Consider a stationary time series defined by

$$X_t = \phi(X_{t-1} + Z_t)$$

where $|\phi| < 1$ and $Z_t \sim WN(0, \sigma^2)$

From formulas 1 and 2

$$P_n X_{n-1} = \mathbf{a}_n' \mathbf{X}_n,$$

where $\mathbf{X}_n = (X_n, \dots, X_1)'$ and

$$\begin{bmatrix} 1 & \phi & \phi^2 & \dots & \phi^{n-1} \\ \phi & 1 & \phi & \dots & \phi^{n-2} \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \phi^{n-1} & \phi^{n-2} & \phi^{n-3} & \dots & 1 \end{bmatrix} \begin{bmatrix} a_1 \\ a_2 \\ \cdot \\ \cdot \\ a_n \end{bmatrix} = \begin{bmatrix} \phi \\ \phi^2 \\ \cdot \\ \cdot \\ \phi^n \end{bmatrix} \quad (*)$$

Example (cont')

A solution of the equation (*) is

$$\mathbf{a}_n = (\phi, 0, \dots, 0)'$$

Mean square error is

$$E(X_{n+h} - P_n X_{n+h})^2 = \sigma^2$$

Exercises

1. Suppose that X_1, X_2, \dots , is a stationary time series with mean μ and ACF $\rho(\cdot)$. Show that the best predictor of X_{n+h} of the form $aX_n + b$ is obtained by choosing $a = \rho(h)$ and $b = \mu(1 - \rho(h))$.
2. Show that the process $X_t = A\cos(\omega t) + B\sin(\omega t), t = 0, \pm 1, \dots$ (where A and B are uncorrelated random variables with mean 0 and variance 1 and ω is a fixed frequency in the interval $[0, \pi]$), is stationary and find its mean and autocovariance function.
3. Find the ACVF of the time series $X_t = Z_t + 0.3Z_{t-1} - 0.4Z_{t-2}$, where $Z_t \sim WN(0, 1)$